### Lecture 1: Brownian motion and stochastic integrals

David Nualart

Department of Mathematics Kansas University

Escuela de Probabilidad CIMAT, Guanajuato, Septiembre 2016

David Nualart (Kansas University)

Sept. 2016 1/45

★ Ξ > ★ Ξ >

- Lecture 1 : Brownian motion, martingales and stochastic integrals.
- 2 Lecture 2 : Introduction to Malliavin calculus.
- Lecture 3 : Stein's method for normal approximations.
- Lecture 4 : Applications to functionals of the fractional Brownian motion.

ヘロア 人間 アメヨア 人口 ア

# Multivariate normal distribution

A random vector X = (X<sub>1</sub>,..., X<sub>n</sub>) has the multivariate normal distribution N(μ, Σ), if its characteristic function is

$$E\left(e^{i\langle u,X
angle}
ight)=\exp\left(i\langle u,\mu
angle-rac{1}{2}u^{T}\Sigma u
ight),\,\,u\in\mathbb{R}^{n},$$

where  $\mu \in \mathbb{R}^n$  and  $\Sigma$  is an  $n \times n$  symmetric and nonnegative definite matrix.

• 
$$\mu = (E(X_1), \ldots, E(X_n))$$

• 
$$\Sigma_{ij} = \operatorname{Cov}(X_i, X_j)$$

If X has the N(μ, Σ) distribution, then Y = AX + b, where A is an m × n matrix and b ∈ ℝ<sup>m</sup>, has the N(Aμ + b, AΣA<sup>T</sup>) distribution.

◆□ ▶ ◆□ ▶ ◆ □ ▶ ◆ □ ▶ ◆ □ ● ● ○ ○ ○

• If  $\Sigma$  is nonsingular, then X has a density given by

$$f(x) = (2\pi)^{-\frac{n}{2}} (\det \Sigma)^{-\frac{1}{2}} \exp\left(-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)\right).$$

Bivariate Normal



ъ

A B + A B +
 A
 B
 A
 B
 A
 B
 A
 B
 A
 B
 A
 B
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A
 A

### Stochastic processes

• A stochastic process  $X = \{X_t, t \ge 0\}$  is a family of random variables

 $X_t: \Omega \to \mathbb{R}$ 

defined on a probability space  $(\Omega, \mathcal{F}, P)$ .

イロト 不得 とくほ とくほとう

### Stochastic processes

• A stochastic process  $X = \{X_t, t \ge 0\}$  is a family of random variables

 $X_t: \Omega \to \mathbb{R}$ 

defined on a probability space  $(\Omega, \mathcal{F}, P)$ .

• The probabilities on  $\mathbb{R}^n$ ,  $n \ge 1$ ,

$$P_{t_1,...,t_n} = P \circ (X_{t_1},...,X_{t_n})^{-1}$$

where  $0 \le t_1 < \cdots < t_n$ , are called the *finite-dimensional marginal distributions* of the process *X*.

◆□ ▶ ◆□ ▶ ◆ □ ▶ ◆ □ ▶ ◆ □ ● ● ○ ○ ○

### Stochastic processes

• A stochastic process  $X = \{X_t, t \ge 0\}$  is a family of random variables

 $X_t: \Omega \to \mathbb{R}$ 

defined on a probability space  $(\Omega, \mathcal{F}, P)$ .

• The probabilities on  $\mathbb{R}^n$ ,  $n \ge 1$ ,

$$P_{t_1,...,t_n} = P \circ (X_{t_1},...,X_{t_n})^{-1}$$

where  $0 \le t_1 < \cdots < t_n$ , are called the *finite-dimensional marginal* distributions of the process *X*.

• For every  $\omega \in \Omega$ , the mapping

$$t \to X_t(\omega)$$

is called a *trajectory* of the process X.

◆□▶ ◆□▶ ◆三▶ ◆三▶ ● ○ ○ ○

### Theorem (Kolmogorov's extension theorem)

Consider a family of probability measures

$$\{P_{t_1,...,t_n}, 0 \le t_1 < \cdots < t_n, n \ge 1\}$$

such that :

- (i)  $P_{t_1,...,t_n}$  is a probability on  $\mathbb{R}^n$ .
- (ii) (Consistency condition) : If  $\{t_{k_1} < \cdots < t_{k_m}\} \subset \{t_1 < \cdots < t_n\}$ , then  $P_{t_{k_1},\ldots,t_{k_m}}$  is the marginal of  $P_{t_1,\ldots,t_n}$ , corresponding to the indexes  $k_1,\ldots,k_m$ .

Then, there exists a stochastic process  $\{X_t, t \ge 0\}$  defined in some probability space  $(\Omega, \mathcal{F}, P)$ , which has the family  $\{P_{t_1,...,t_n}\}$  as finite-dimensional marginal distributions.

Take Ω as the set of all functions ω : [0,∞) → ℝ, F the σ-algebra generated by cylindrical sets, extend the probability from cylindrical sets to F, and set X<sub>t</sub>(ω) = ω(t).

◆□▶ ◆□▶ ◆三▶ ◆三▶ ● ○ ○ ○

# Gaussian processes

X = {X<sub>t</sub>, t ≥ 0} is called *Gaussian* if all its finite-dimensional marginal distributions are multivariate normal.

イロン 不同 とくほ とくほ とう

# Gaussian processes

- X = {X<sub>t</sub>, t ≥ 0} is called *Gaussian* if all its finite-dimensional marginal distributions are multivariate normal.
- The law of a Gaussian process is determined by the mean function  $E(X_t)$  and the covariance function

$$\operatorname{Cov}(X_t, X_s) = E((X_t - E(X_t))(X_s - E(X_s))).$$

ヘロン ヘアン ヘビン ヘビン

# Gaussian processes

- X = {X<sub>t</sub>, t ≥ 0} is called *Gaussian* if all its finite-dimensional marginal distributions are multivariate normal.
- The law of a Gaussian process is determined by the mean function  $E(X_t)$  and the covariance function

$$\operatorname{Cov}(X_t, X_s) = E((X_t - E(X_t))(X_s - E(X_s))).$$

Suppose μ : ℝ<sub>+</sub> → ℝ, and Γ : ℝ<sub>+</sub> × ℝ<sub>+</sub> → ℝ is symmetric and nonnegative definite :

$$\sum_{i,j=1}^n \Gamma(t_i,t_j) a_i a_j \geq 0, \quad \forall \ t_i \geq 0, \ a_i \in \mathbb{R}.$$

Then there exists a Gaussian process with mean  $\mu$  and covariance function  $\Gamma$ .

・ロト ・ 同ト ・ ヨト ・ ヨト … ヨ

### Equivalent processes

Two processes, X, Y are *equivalent* (or X is a version of Y) if for all t ≥ 0,

$$P\{X_t = Y_t\} = 1$$

### Equivalent processes

Two processes, X, Y are equivalent (or X is a version of Y) if for all t ≥ 0,

$$P\{X_t = Y_t\} = 1.$$

 Two equivalent processes may have quite different trajectories. For example, the processes X<sub>t</sub> = 0 for all t ≥ 0 and

$$Y_t = \begin{cases} 0 & \text{if } \xi \neq t \\ 1 & \text{if } \xi = t \end{cases}$$

where  $\xi \ge 0$  is a continuous random variable, are equivalent, because  $P(\xi = t) = 0$ , but their trajectories are different.

・ロト ・ 理 ト ・ ヨ ト ・

# Equivalent processes

Two processes, X, Y are equivalent (or X is a version of Y) if for all t ≥ 0,

$$P\{X_t = Y_t\} = 1.$$

 Two equivalent processes may have quite different trajectories. For example, the processes X<sub>t</sub> = 0 for all t ≥ 0 and

$$Y_t = \begin{cases} 0 & \text{if } \xi \neq t \\ 1 & \text{if } \xi = t \end{cases}$$

where  $\xi \ge 0$  is a continuous random variable, are equivalent, because  $P(\xi = t) = 0$ , but their trajectories are different.

• Two processes X and Y are said to be *indistinguishable* if

$$X_t(\omega) = Y_t(\omega)$$

for all  $t \ge 0$  and for all  $\omega \in \Omega^*$ , with  $P(\Omega^*) = 1$ .

 Two equivalent processes with right-continuous trajectories are indistinguishable.

・ロト ・ 理 ト ・ ヨ ト ・

Theorem (Kolmogorov's continuity theorem)

Suppose that  $X = \{X_t, t \in [0, T]\}$  satisfies

$$E(|X_t - X_s|^{\beta}) \leq K|t - s|^{1+\alpha},$$

for all  $s, t \in [0, T]$ , and for some constants  $\beta, \alpha > 0$ . Then, there exists a version  $\widetilde{X}$  of X such that, if  $\gamma < \alpha/\beta$ ,

$$|\widetilde{X}_t - \widetilde{X}_{m{s}}| \leq G_\gamma |t - m{s}|^\gamma$$

for all  $s, t \in [0, T]$ , where  $G_{\gamma}$  is a random variable.

• The trajectories of  $\widetilde{X}$  are Hölder continuous of order  $\gamma$  for any  $\gamma < \alpha/\beta$ .

・ロ・ ・ 同・ ・ ヨ・ ・ ヨ・

A stochastic process  $B = \{B_t, t \ge 0\}$  is called a *Brownian motion* if :

- i)  $B_0 = 0$  almost surely.
- ii) Independent increments : For all  $0 \le t_1 < \cdots < t_n$  the increments  $B_{t_n} B_{t_{n-1}}, \ldots, B_{t_2} B_{t_1}$ , are independent random variables.
- iii) If  $0 \le s < t$ , the increment  $B_t B_s$  has the normal distribution N(0, t s).
- iv) With probability one,  $t \rightarrow B_t(\omega)$  is continuous.

◆□▶ ◆□▶ ◆三▶ ◆三▶ ● ○ ○ ○

### Proposition

Properties i), ii), iii) are equivalent to :

 $(\star)$  B is a Gaussian process with mean zero and covariance

 $\Gamma(\boldsymbol{s},t)=\min(\boldsymbol{s},t).$ 

### Proposition

Properties i), ii), iii) are equivalent to :

(\*) B is a Gaussian process with mean zero and covariance

 $\Gamma(s, t) = \min(s, t).$ 

Proof :

a) Suppose i), i) and iii). The distribution of  $(B_{t_1}, \ldots, B_{t_n})$ , for  $0 < t_1 < \cdots < t_n$ , is normal, because this vector is a linear transformation of  $(B_{t_1}, B_{t_2} - B_{t_1}, \ldots, B_{t_n} - B_{t_{n-1}})$  which has independent and normal components.

The mean is zero, and for s < t, the covariance is

$$E(B_sB_t) = E(B_s(B_t - B_s + B_s)) = E(B_s(B_t - B_s)) + E(B_s^2) = s.$$

b) The converse is also easy to show.  $\Box$ 

・ロト ・ 同ト ・ ヨト ・ ヨト … ヨ

# Construction of the Brownian motion

1. The function  $\Gamma(s, t) = \min(s, t)$  is symmetric and nonnegative definite because it can be written as

$$\min(s,t) = \int_0^\infty \mathbf{1}_{[0,s]}(r) \mathbf{1}_{[0,t]}(r) dr,$$

S0

$$\sum_{i,j=1}^{n} a_{i}a_{j}\min(t_{i},t_{j}) = \sum_{i,j=1}^{n} a_{i}a_{j}\int_{0}^{\infty} \mathbf{1}_{[0,t_{i}]}(r)\mathbf{1}_{[0,t_{j}]}(r)dr$$
$$= \int_{0}^{\infty} \left[\sum_{i=1}^{n} a_{i}\mathbf{1}_{[0,t_{i}]}(r)\right]^{2}dr \ge 0.$$

Therefore, by Kolmogorov's extension theorem there exists a Gaussian process *B* with zero mean and covariance function min(s, t).

ヘロト ヘワト ヘビト ヘビト

#### 2. The process B satisfies

$$E\left[\left(B_{t}-B_{s}\right)^{2k}
ight]=rac{(2k)!}{2^{k}k!}(t-s)^{k}, \quad s\leq t$$

for any  $k \ge 1$ , because the distribution of  $B_t - B_s$  is N(0, t - s).

**3.** Therefore, by the Kolmogorov's continuity theorem, there exist a version  $\widetilde{B}$  of *B*, such that  $\widetilde{B}$  has Hölder continuous trajectories of order  $\gamma$  for any  $\gamma < \frac{k-1}{2k}$  on any interval [0, T]. This implies that the paths are  $\gamma$ -Hölder on [0, T] for any  $\gamma < \frac{1}{2}$  and for any T > 0.

◆□▶ ◆□▶ ◆三▶ ◆三▶ ● ○ ○ ○

# Brownian motion and random walk

- Let {*ξ<sub>k</sub>*, 1 ≤ *k* ≤ *n*} be independent and identically distributed random variables with zero mean and variance one.
- Define  $S_n(0) = 0$ ,

$$S_n(\frac{kT}{n}) = \sqrt{T}\frac{\xi_1 + \dots + \xi_k}{\sqrt{n}}, \quad k = 1, \dots, n$$

and extend  $S_n(t)$  to  $t \in [0, T]$  by linear interpolation.

• Donsker Invariance Principle : The law of the random walk  $S_n$  on C([0, T]) converges to the Wiener measure, which is the law of the Brownian motion. That is, that for any continuous and bounded function  $\varphi : C([0, T]) \to \mathbb{R}$ ,

$$E(\varphi(S_n)) \stackrel{n \to \infty}{\longrightarrow} E(\varphi(B)),$$

◆□ ▶ ◆□ ▶ ◆ □ ▶ ◆ □ ▶ ◆ □ ● ● ○ ○ ○

### Simulations of Brownian motion



Sept. 2016 15/45

# **Basic properties**

1. Selfsimilarity :

For any a > 0, the process  $\{a^{-\frac{1}{2}}B_{at}, t \ge 0\}$  is also a Brownian motion.



#### 🖹 🔊 ९ ( ભ

- 2. For any h > 0, the process  $\{B_{t+h} B_h, t \ge 0\}$  is a Brownian motion.
- 3. The process  $\{-B_t, t \ge 0\}$  is a Brownian motion.
- 4. The process

$$X_t = \begin{cases} tB_{1/t}, & t > 0\\ 0, & t = 0 \end{cases}$$

is a Brownian motion.

◆□ ▶ ◆□ ▶ ◆ □ ▶ ◆ □ ▶ ● □ ● ● ● ●

### Quadratic variation

Fix a time interval [0, t] and consider a partition

$$\pi = \{ \mathbf{0} = t_0 < t_1 < \cdots < t_n = t \}.$$

Define  $\Delta t_k = t_k - t_{k-1}$ ,  $\Delta B_k = B_{t_k} - B_{t_{k-1}}$  and  $|\pi| = \max_{1 \le k \le n} \Delta t_k$ .

### Proposition

The following convergence holds in  $L^2$ :

$$\lim_{\pi|\to 0}\sum_{k=1}^n \left(\Delta B_k\right)^2 = t.$$

• We can say that  $(\Delta B_t)^2 \sim \Delta t$ 

David Nualart (Kansas University)

Sept. 2016 18/45

◆□ ▶ ◆□ ▶ ◆ □ ▶ ◆ □ ▶ ◆ □ ● ● ○ ○ ○

*Proof :* Set  $\xi_k = (\Delta B_k)^2 - \Delta t_k$ . The random variables  $\xi_k$  are independent and centered. Thus,

$$E\left[\left(\sum_{k=1}^{n} (\Delta B_{k})^{2} - t\right)^{2}\right] = E\left[\left(\sum_{k=1}^{n} \xi_{k}\right)^{2}\right] = \sum_{k=1}^{n} E\left[\xi_{k}^{2}\right]$$
$$= \sum_{k=1}^{n} \left[3\left(\Delta t_{k}\right)^{2} - 2\left(\Delta t_{k}\right)^{2} + \left(\Delta t_{k}\right)^{2}\right]$$
$$= 2\sum_{k=1}^{n} (\Delta t_{k})^{2} \le 2t|\pi| \xrightarrow{|\pi| \to 0} 0. \quad \Box$$

Sept. 2016 19/45

*Proof :* Set  $\xi_k = (\Delta B_k)^2 - \Delta t_k$ . The random variables  $\xi_k$  are independent and centered. Thus,

$$E\left[\left(\sum_{k=1}^{n} (\Delta B_k)^2 - t\right)^2\right] = E\left[\left(\sum_{k=1}^{n} \xi_k\right)^2\right] = \sum_{k=1}^{n} E\left[\xi_k^2\right]$$
$$= \sum_{k=1}^{n} \left[3\left(\Delta t_k\right)^2 - 2\left(\Delta t_k\right)^2 + \left(\Delta t_k\right)^2\right]$$
$$= 2\sum_{k=1}^{n} (\Delta t_k)^2 \le 2t|\pi| \xrightarrow{|\pi| \to 0} 0. \quad \Box$$

*Exercise* : Using the Borel-Cantelli lemma, show that if  $\{\pi^n\}$  is a sequence of partitions of [0, t] such that  $\sum_n |\pi^n| < \infty$ , then  $\sum_{k=1}^n (\Delta B_k)^2$  converges almost surely to *t*.

◆□ ▶ ◆□ ▶ ◆ □ ▶ ◆ □ ▶ ● □ ● ● ● ●

### Infinite total variation

Define

$$V_t = \sup_{\pi} \sum_{k=1}^n |\Delta B_k|$$

Then,

$$P(V_t = \infty) = 1.$$

In fact, using the continuity of the trajectories of the Brownian motion, we have, on the set  $V < \infty$ ,

$$\sum_{k=1}^{n} \left( \Delta B_k \right)^2 \leq \sup_k \left| \Delta B_k \right| \left( \sum_{k=1}^{n} \left| \Delta B_k \right| \right) \leq V \sup_k \left| \Delta B_k \right| \stackrel{|\pi| \to 0}{\longrightarrow} 0.$$

Then,  $V < \infty$  contradicts the fact that  $\sum_{k=1}^{n} (\Delta B_k)^2$  converges in  $L^2$  to t as  $|\pi| \to 0$ .

Sept. 2016 20/45

= 900

イロト 不得 とくほ とくほ とう

# Martingales

We assume that {*F<sub>t</sub>*, *t* ≥ 0} is an increasing family of *σ*-fields, contained in *F* (*filtration*).

### Definition

An adapted process  $M = \{M_t, t \ge 0\}$  is called a *martingale* with respect to  $\mathcal{F}_t$  if

- (i) For all  $t \ge 0$ ,  $E(|M_t|) < \infty$ .
- (ii) For each  $s \leq t$ ,  $E(M_t | \mathcal{F}_s) = M_s$ .

# Martingales

We assume that {*F<sub>t</sub>*, *t* ≥ 0} is an increasing family of *σ*-fields, contained in *F* (*filtration*).

### Definition

An adapted process  $M = \{M_t, t \ge 0\}$  is called a *martingale* with respect to  $\mathcal{F}_t$  if

(i) For all 
$$t \ge 0$$
,  $E(|M_t|) < \infty$ .

(ii) For each 
$$s \leq t$$
,  $E(M_t | \mathcal{F}_s) = M_s$ .

Property (ii) can also be written as :

$$E(M_t - M_s | \mathcal{F}_s) = 0.$$

David Nualart (Kansas University)

Let  $B_t$  be a Brownian motion and let  $\mathcal{F}_t$  be the filtration generated by  $B_t$ :

$$\mathcal{F}_t = \sigma\{\boldsymbol{B}_s, \boldsymbol{0} \leq \boldsymbol{s} \leq t\}.$$

Then, the processes

$$M_t^{(1)} = B_t$$
  

$$M_t^{(2)} = B_t^2 - t$$
  

$$M_t^{(3)} = \exp(aB_t - \frac{a^2t}{2})$$

where  $a \in \mathbb{R}$ , are martingales.

1.  $B_t$  is a martingale because

$$E(B_t-B_s|\mathcal{F}_s)=E(B_t-B_s)=0.$$

・ロト・雪・・ヨト・ヨー めんの

David Nualart (Kansas University)

Sept. 2016 23/45

1.  $B_t$  is a martingale because

$$E(B_t - B_s | \mathcal{F}_s) = E(B_t - B_s) = 0$$

2. For  $B_t^2 - t$ , we can write, using the properties of the conditional expectation, for s < t

$$\begin{split} E(B_t^2 | \mathcal{F}_s) &= E((B_t - B_s + B_s)^2 | \mathcal{F}_s) \\ &= E((B_t - B_s )^2 | \mathcal{F}_s) + 2E((B_t - B_s ) B_s | \mathcal{F}_s) \\ &+ E(B_s^2 | \mathcal{F}_s) \\ &= E(B_t - B_s )^2 + 2B_s E((B_t - B_s ) | \mathcal{F}_s) + B_s^2 \\ &= t - s + B_s^2. \end{split}$$

Sept. 2016 23/45

イロン 不同 とくほ とくほ とう

1.  $B_t$  is a martingale because

$$E(B_t - B_s | \mathcal{F}_s) = E(B_t - B_s) = 0$$

2. For  $B_t^2 - t$ , we can write, using the properties of the conditional expectation, for s < t

$$E(B_t^2 | \mathcal{F}_s) = E((B_t - B_s + B_s)^2 | \mathcal{F}_s)$$
  
=  $E((B_t - B_s)^2 | \mathcal{F}_s) + 2E((B_t - B_s) | B_s | \mathcal{F}_s)$   
 $+ E(B_s^2 | \mathcal{F}_s)$   
=  $E(B_t - B_s)^2 + 2B_s E((B_t - B_s) | \mathcal{F}_s) + B_s^2$   
=  $t - s + B_s^2$ .

3. Finally, for  $\exp(aB_t - \frac{a^2t}{2})$  we have  $E(e^{aB_t - \frac{a^2t}{2}} | \mathcal{F}_s) = e^{aB_s} E(e^{a(B_t - B_s) - \frac{a^2t}{2}} | \mathcal{F}_s)$   $= e^{aB_s} E(e^{a(B_t - B_s) - \frac{a^2t}{2}})$  $= e^{aB_s} e^{\frac{a^2(t-s)}{2} - \frac{a^2t}{2}} = e^{aB_s - \frac{a^2s}{2}}$ 

### Theorem

Let  $\{M_t, t \in [0, T]\}$  be a continuous martingale such that  $E(|M_T|^p) < \infty$  for some  $p \ge 1$ . Then, for all  $\lambda > 0$  we have

$$P\left(\sup_{0\leq t\leq T}|M_t|>\lambda\right)\leq \frac{1}{\lambda^p}E(|M_T|^p).$$
(1)

If p > 1, then

$$E\left(\sup_{0\leq t\leq T}|M_t|^p\right)\leq \left(\frac{p}{p-1}\right)^p E(|M_T|^p).$$
(2)

◆□ ▶ ◆□ ▶ ◆ □ ▶ ◆ □ ▶ ◆ □ ● ● ○ ○ ○

### The Wiener integral

• The integral of a step function  $\varphi_t = \sum_{j=0}^{m-1} a_j \mathbf{1}_{(t_j, t_{j+1}]}(t) \in \mathcal{E}$  with respect to a Brownian motion B on [0, T] is defined by

$$\int_{0}^{T} \varphi_{t} dB_{t} = \sum_{j=0}^{m-1} a_{j} (B_{t_{j+1}} - B_{t_{j}})$$
### The Wiener integral

• The integral of a step function  $\varphi_t = \sum_{j=0}^{m-1} a_j \mathbf{1}_{(t_j, t_{j+1}]}(t) \in \mathcal{E}$  with respect to a Brownian motion B on [0, T] is defined by

$$\int_0^T \varphi_t dB_t = \sum_{j=0}^{m-1} a_j (B_{t_{j+1}} - B_{t_j})$$

The mapping φ → ∫<sub>0</sub><sup>T</sup> φ<sub>t</sub>dB<sub>t</sub> from ε ⊂ L<sup>2</sup>([0, T]) to L<sup>2</sup>(Ω) is linear and isometric :

$$E\left[\left(\int_{0}^{T}\varphi_{t}dB_{t}\right)^{2}\right]=\sum_{j=0}^{m-1}a_{j}^{2}(t_{j+1}-t_{j})=\int_{0}^{T}\varphi_{t}^{2}dt=\|\varphi\|_{L^{2}([0,T])}^{2}.$$

◆□ ▶ ◆□ ▶ ◆ □ ▶ ◆ □ ▶ ◆ □ ● ● ○ ○ ○

# The Wiener integral

The integral of a step function φ<sub>t</sub> = ∑<sub>j=0</sub><sup>m-1</sup> a<sub>j</sub> 1<sub>(t<sub>j</sub>,t<sub>j+1</sub>]</sub>(t) ∈ ε with respect to a Brownian motion B on [0, T] is defined by

$$\int_0^T \varphi_t dB_t = \sum_{j=0}^{m-1} a_j (B_{t_{j+1}} - B_{t_j})$$

The mapping φ → ∫<sub>0</sub><sup>T</sup> φ<sub>t</sub> dB<sub>t</sub> from ε ⊂ L<sup>2</sup>([0, T]) to L<sup>2</sup>(Ω) is linear and isometric :

$$E\left[\left(\int_{0}^{T}\varphi_{t}dB_{t}\right)^{2}\right]=\sum_{j=0}^{m-1}a_{j}^{2}(t_{j+1}-t_{j})=\int_{0}^{T}\varphi_{t}^{2}dt=\|\varphi\|_{L^{2}([0,T])}^{2}.$$

•  $\mathcal{E}$  is a dense subspace of  $L^2([0, T])$ . Therefore, the mapping

$$\varphi \to B(\varphi) =: \int_0^T \varphi_t dB_t$$

can be extended to a linear isometry between  $L^2([0, T])$  and the Gaussian subspace of  $L^2(\Omega)$  spanned by  $\{B_t, t \in [0, T]\}$ .

David Nualart (Kansas University)

Sept. 2016 25/45

# Progressively measurable processes

Let  $\mathcal{F}_t$  be the  $\sigma$ -field generated by the random variables  $\{B_s, 0 \le s \le t\}$  and the sets of probability zero.

### Definition

We say that  $u = \{u_t, t \in [0, T]\}$  is *progressively measurable* if for any  $t \in [0, T]$ , the restriction of u to  $\Omega \times [0, t]$  is  $\mathcal{F}_t \times \mathcal{B}([0, t])$ -measurable.

イロト イヨト イヨト イ

# Progressively measurable processes

Let  $\mathcal{F}_t$  be the  $\sigma$ -field generated by the random variables  $\{B_s, 0 \le s \le t\}$  and the sets of probability zero.

### Definition

We say that  $u = \{u_t, t \in [0, T]\}$  is *progressively measurable* if for any  $t \in [0, T]$ , the restriction of u to  $\Omega \times [0, t]$  is  $\mathcal{F}_t \times \mathcal{B}([0, t])$ -measurable.

- Let P be the σ-field of sets A ⊂ Ω × [0, T] such that 1<sub>A</sub> is progressively measurable.
- We denote by L<sup>2</sup><sub>T</sub>(P) the Hilbert space L<sup>2</sup>(Ω × [0, T], P, P × ℓ), where ℓ is the Lebesgue measure, equipped with the norm

$$\|u\|^2 = E\left(\int_0^T u_s^2 ds\right).$$

ヘロト ヘワト ヘビト ヘビト

### Stochastic integrals

•  $u = \{u_t, t \in [0, T]\}$  is a simple process if

$$u_t = \sum_{j=0}^{n-1} \phi_j \mathbf{1}_{(t_j, t_{j+1}]}(t),$$

where  $0 \le t_0 \le t_1 \le \cdots \le t_n = T$  and  $\phi_j$  are  $\mathcal{F}_{t_j}$ -measurable random variables such that  $E(\phi_j^2) < \infty$ .

◆□ ▶ ◆□ ▶ ◆ □ ▶ ◆ □ ▶ ◆ □ ● ● ○ ○ ○

### Stochastic integrals

•  $u = \{u_t, t \in [0, T]\}$  is a simple process if

$$u_t = \sum_{j=0}^{n-1} \phi_j \mathbf{1}_{(t_j, t_{j+1}]}(t),$$

where  $0 \le t_0 \le t_1 \le \cdots \le t_n = T$  and  $\phi_j$  are  $\mathcal{F}_{t_j}$ -measurable random variables such that  $E(\phi_j^2) < \infty$ .

• We define the stochastic integral of *u* as

$$I(u) := \int_0^T u_t dB_t = \sum_{j=0}^{n-1} \phi_j \left( B_{t_{j+1}} - B_{t_j} \right).$$

◆□ ▶ ◆□ ▶ ◆ □ ▶ ◆ □ ▶ ◆ □ ● ● ○ ○ ○

#### Properties of the stochastic integral of simple processes

(i) Linearity :

$$\int_0^T (au_t + bv_t) dB_t = a \int_0^T u_t dB_t + b \int_0^T v_t dB_t.$$

(ii) Zero mean :

$$E\left(\int_0^T u_t dB_t\right) = 0.$$

In fact,

$$E\left(\int_{0}^{T} u_{t} dB_{t}\right) = \sum_{j=0}^{n-1} E\left[\phi_{j}\left(B_{t_{j+1}} - B_{t_{j}}\right)\right]$$
$$= \sum_{j=0}^{n-1} E[\phi_{j}]E[B_{t_{j+1}} - B_{t_{j}}] = 0.$$

(iii) Isometry property :

$$E\left[\left(\int_0^T u_t dB_t\right)^2\right] = E\left(\int_0^T u_t^2 dt\right).$$

トレット 山田 マール・ 山田 マート・

David Nualart (Kansas University)

Sept. 2016 29/45

(iii) Isometry property :

$$E\left[\left(\int_0^T u_t dB_t\right)^2\right] = E\left(\int_0^T u_t^2 dt\right).$$

*Proof :* Set  $\Delta B_j = B_{t_{j+1}} - B_{t_j}$ . Then

$$E\left(\phi_{i}\phi_{j}\Delta B_{i}\Delta B_{j}\right) = \begin{cases} 0 & \text{if } i \neq j \\ E\left(\phi_{j}^{2}\right)\left(t_{j+1} - t_{j}\right) & \text{if } i = j \end{cases}$$

because if i < j the random variables  $\phi_i \phi_j \Delta B_i$  and  $\Delta B_j$  are independent and if i = j the random variables  $\phi_i^2$  and  $(\Delta B_i)^2$  are independent. So, we obtain

$$E\left[\left(\int_{0}^{T} u_{t} dB_{t}\right)^{2}\right] = \sum_{i,j=0}^{n-1} E\left(\phi_{i} \phi_{j} \Delta B_{i} \Delta B_{j}\right) = \sum_{i=0}^{n-1} E\left(\phi_{i}^{2}\right)\left(t_{i+1} - t_{i}\right)$$
$$= E\left(\int_{0}^{T} u_{t}^{2} dt\right). \quad \Box$$

◆□ ▶ ◆□ ▶ ◆ □ ▶ ◆ □ ▶ ◆ □ ● ● ○ ○ ○

### Proposition

The space  $\mathcal{E}$  of simple processes is dense in  $L^2_T(\mathcal{P})$ .

David Nualart (Kansas University)

### Proposition

The space  $\mathcal{E}$  of simple processes is dense in  $L^2_T(\mathcal{P})$ .

# *Proof :* Use the approximating sequence

$$u_t^{(n)} = \sum_{j=1}^{n-1} \left( \frac{n}{T} \int_{t_{j-1}}^{t_j} u_s ds \right) \mathbf{1}_{(t_j, t_{j+1}]}(t),$$

where  $t_j = \frac{jT}{n}$ .  $\Box$ 

ヘロト ヘワト ヘビト ヘビト

### Proposition

The stochastic integral can be extended to a linear isometry :

 $I: L^2_T(\mathcal{P}) \to L^2(\Omega).$ 

*Proof :* This follows form the fact that  $\mathcal{E}$  is dense in  $L^2_T(\mathcal{P})$ .  $\Box$ .

• The stochastic integral has the following properties :

 $E\left[I(u)\right]=0$ 

and

$$E[I(u)I(v)] = E\left(\int_0^\infty u_s v_s ds\right)$$

イロン 不得 とくほ とくほ とうほ

$$\int_0^T B_t dB_t = \frac{1}{2}B_T^2 - \frac{1}{2}T$$

*Proof :* The process  $B_t$  being continuous in mean square, we can choose as approximating sequence

$$u_t^{(n)} = \sum_{j=1}^n B_{t_{j-1}} \mathbf{1}_{(t_{j-1},t_j]}(t),$$

l

where  $t_j = \frac{jT}{n}$ , and we obtain

$$\int_{0}^{T} B_{t} dB_{t} = \lim_{n \to \infty} \sum_{j=1}^{n} B_{t_{j-1}} \left( B_{t_{j}} - B_{t_{j-1}} \right)$$
$$= \frac{1}{2} \lim_{n \to \infty} \sum_{j=1}^{n} \left( B_{t_{j}}^{2} - B_{t_{j-1}}^{2} \right) - \frac{1}{2} \lim_{n \to \infty} \sum_{j=1}^{n} \left( B_{t_{j}} - B_{t_{j-1}} \right)^{2}$$
$$= \frac{1}{2} B_{T}^{2} - \frac{1}{2} T.$$

David Nualart (Kansas University)

Sept. 2016 32/45

# Indefinite stochastic integrals

For  $u \in L^2_T(\mathcal{P})$ , we define the stochastic process

$$\int_0^t u_s dB_s := \int_0^T u_s \mathbf{1}_{[0,t]}(s) dB_s , \quad t \in [0,T]$$

#### Proposition

Let  $u \in L^2_T(\mathcal{P})$ . The indefinite stochastic integral

$$M_t = \int_0^t u_s dB_s$$

is a square integrable martingale with respect to the filtration  $\mathcal{F}_t$  and admits a continuous version.

・ロト ・ 同ト ・ ヨト ・ ヨト … ヨ

# ltô's formula

- Itô's stochastic integral does not follow the chain rule of classical calculus.
- Example :

$$\int_0^t B_s dB_s = \frac{1}{2}B_t^2 - \frac{t}{2},$$

whereas if  $x_t$  is a differentiable function such that  $x_0 = 0$ ,

$$\int_0^t x_s dx_s = \int_0^t x_s x_s' ds = \frac{1}{2} x_t^2.$$

In differential form

$$d(B_t^2)=2B_tdB_t+dt,$$

and *dt* comes from  $(dB_t)^2 \sim dt$  and the Taylor expansion up to the second order.

イロト 不得 とくほ とくほとう

- The stochastic integral can be extended (using convergence in probability) to progressively measurable processes satisfying  $\int_0^T u_s^2 ds < \infty$  a.s. Denote the class of those processes by  $L^2_{T,loc}(\mathcal{P})$ .
- Denote by  $L^{1}_{T,loc}(\mathcal{P})$  the space of progressively measurable processes  $v = \{v_t, t \in [0, T]\}$  such that for  $\int_0^T |v_s| ds < \infty$  a.s.

#### Theorem (Itô's formula)

Suppose that

$$X_t = X_0 + \int_0^t u_s dB_s + \int_0^t v_s ds,$$

where  $u \in L^2_{T,loc}(\mathcal{P})$  and  $v \in L^1_{T,loc}(\mathcal{P})$ . Let  $f \in C^{1,2}$ . Then,

$$Y_t = f(0, X_0) + \int_0^t \frac{\partial f}{\partial t}(s, X_s) ds + \int_0^t \frac{\partial f}{\partial x}(s, X_s) u_s dB_s \\ + \int_0^t \frac{\partial f}{\partial x}(s, X_s) v_s ds + \frac{1}{2} \int_0^t \frac{\partial^2 f}{\partial x^2}(s, X_s) u_s^2 ds.$$

In differential notation Itô's formula can be written as

$$df(t, X_t) = \frac{\partial f}{\partial t}(t, X_t)dt + \frac{\partial f}{\partial x}(t, X_t)dX_t + \frac{1}{2}\frac{\partial^2 f}{\partial x^2}(s, X_s)(dX_t)^2,$$

where  $(dX_t)^2$  is computed from

$$dX_t = u_t dB_t + v_t dt,$$

using the product rule

×	$dB_t$	dt
$dB_t$	dt	0
dt	0	0

# Multiple stochastic integrals

*L*<sup>2</sup><sub>s</sub>([0, *T*]<sup>n</sup>) is the space of symmetric square integrable functions
 *f*: [0, *T*]<sup>n</sup> → ℝ.

# Multiple stochastic integrals

- L<sup>2</sup><sub>s</sub>([0, *T*]<sup>n</sup>) is the space of symmetric square integrable functions *f* : [0, *T*]<sup>n</sup> → ℝ.
- For any  $f \in L^2_s([0, T]^n)$

$$||f||^2_{L^2([0,T]^n)} = n! \int_{\Delta_n} f^2(t_1,\ldots,t_n) dt_1 \cdots dt_n,$$

where

$$\Delta_n = \{ (t_1, \ldots, t_n) \in [0, T]^n : 0 < t_1 < \cdots < t_n < T \}.$$

# Multiple stochastic integrals

- $L^2_s([0, T]^n)$  is the space of symmetric square integrable functions  $f: [0, T]^n \to \mathbb{R}$ .
- For any  $f \in L^2_s([0, T]^n)$

$$||f||^2_{L^2([0,T]^n)} = n! \int_{\Delta_n} f^2(t_1,\ldots,t_n) dt_1 \cdots dt_n,$$

where

$$\Delta_n = \{ (t_1, \ldots, t_n) \in [0, T]^n : 0 < t_1 < \cdots < t_n < T \}.$$

• If  $f : [0, T]^n \to \mathbb{R}$  we define its symmetrization as

$$\widetilde{f}(t_1,\ldots,t_n)=\frac{1}{n!}\sum_{\sigma}f(t_{\sigma(1)},\ldots,t_{\sigma(n)}),$$

where the sum runs over all permutations  $\sigma$  of  $\{1, 2, ..., n\}$ .

Sept. 2016 37/45

◆□▶ ◆□▶ ◆三▶ ◆三▶ ● ○ ○ ○

The multiple stochastic integral of f ∈ L<sup>2</sup><sub>s</sub>([0, T]<sup>n</sup>) is defined as an iterated Itô integral :

$$I_n(f) = n! \int_0^T \int_0^{t_n} \cdots \int_0^{t_2} f(t_1, \ldots, t_n) dB_{t_1} \cdots dB_{t_n}.$$

Sept. 2016 38/45

The multiple stochastic integral of f ∈ L<sup>2</sup><sub>s</sub>([0, T]<sup>n</sup>) is defined as an iterated Itô integral :

$$I_n(f) = n! \int_0^T \int_0^{t_n} \cdots \int_0^{t_2} f(t_1, \ldots, t_n) dB_{t_1} \cdots dB_{t_n}.$$

• We have the following property :

$$E[I_n(f)I_m(g)] = \begin{cases} 0 & \text{if } n \neq m \\ n! \langle f, g \rangle_{L^2([0,T]^n)} & \text{if } n = m. \end{cases}$$

The multiple stochastic integral of f ∈ L<sup>2</sup><sub>s</sub>([0, T]<sup>n</sup>) is defined as an iterated Itô integral :

$$I_n(f) = n! \int_0^T \int_0^{t_n} \cdots \int_0^{t_2} f(t_1, \ldots, t_n) dB_{t_1} \cdots dB_{t_n}.$$

• We have the following property :

$$E[I_n(f)I_m(g)] = \begin{cases} 0 & \text{if } n \neq m \\ n! \langle f, g \rangle_{L^2([0,T]^n)} & \text{if } n = m. \end{cases}$$

If *f* ∈ *L*<sup>2</sup>([0, *T*]<sup>*n*</sup>) is not necessarily symmetric we define
 *I<sub>n</sub>*(*f*) = *I<sub>n</sub>*(*t̃*).

◆□ ▶ ◆□ ▶ ◆ □ ▶ ◆ □ ▶ ◆ □ ● ● ○ ○ ○

$$h_n(x) = (-1)^n e^{x^2/2} \frac{d^n}{dx^n} (e^{-x^2/2}), \quad n \ge 1.$$

Sept. 2016 39/45

$$h_n(x) = (-1)^n e^{x^2/2} \frac{d^n}{dx^n} (e^{-x^2/2}), \quad n \ge 1.$$

Elementary properties :

$$h'_n(x) = nh_{n-1}(x)$$
  
 $h_{n+1}(x) = xh_n(x) - h'_n(x) = xh_n(x) - nh_{n-1}(x).$ 

∃ 9900

・ロト ・聞ト ・ヨト ・ヨト

$$h_n(x) = (-1)^n e^{x^2/2} \frac{d^n}{dx^n} (e^{-x^2/2}), \quad n \ge 1.$$

Elementary properties :

$$h'_n(x) = nh_{n-1}(x)$$
  
 $h_{n+1}(x) = xh_n(x) - h'_n(x) = xh_n(x) - nh_{n-1}(x).$ 

• The first Hermite poynomials are  $h_1(x) = x$ ,  $h_2(x) = x^2 - 1$ ,  $h_3(x) = x^3 - 3x$ , ....

▲ロト ▲帰 ト ▲ 臣 ト ▲ 臣 ト ○ 臣 - の Q ()

$$h_n(x) = (-1)^n e^{x^2/2} \frac{d^n}{dx^n} (e^{-x^2/2}), \quad n \ge 1.$$

Elementary properties :

$$h'_n(x) = nh_{n-1}(x)$$
  
 $h_{n+1}(x) = xh_n(x) - h'_n(x) = xh_n(x) - nh_{n-1}(x).$ 

- The first Hermite poynomials are  $h_1(x) = x$ ,  $h_2(x) = x^2 1$ ,  $h_3(x) = x^3 3x$ , ....
- For any  $a \in \mathbb{R}$ ,

$$e^{az-\frac{1}{2}a^2} = \sum_{n=0}^{\infty} \frac{a^n}{n!} h_n(z).$$

◆□ ▶ ◆□ ▶ ◆ □ ▶ ◆ □ ▶ ◆ □ ● ● ○ ○ ○

### Theorem

### For any $g \in L^2([0, T])$ such that $\|g\|_{L^2([0, T])} = 1$ , we have

$$I_n(g^{\otimes n}) = h_n\left(\int_0^T g_t dB_t\right)$$

where  $g^{\otimes n}(t_1,\ldots,t_n) = g(t_1)\cdots g(t_n)$ .

David Nualart (Kansas University)

Sept. 2016 40/45

◆□ ▶ ◆□ ▶ ◆ □ ▶ ◆ □ ▶ ◆ □ ● ● ○ ○ ○

#### Theorem

For any  $g \in L^2([0, T])$  such that  $\|g\|_{L^2([0, T])} = 1$ , we have

$$I_n(g^{\otimes n}) = h_n\left(\int_0^T g_t dB_t\right)$$

where  $g^{\otimes n}(t_1,\ldots,t_n) = g(t_1)\cdots g(t_n)$ .

Proof :

(i) Fix  $a \in \mathbb{R}$  and set

$$M_t = \exp\left(a\int_0^t g_s dB_s - \frac{1}{2}a^2\int_0^t g_s^2 ds
ight).$$

One one hand, we have

$$M_T = e^{a \int_0^T g_s dB_s - \frac{1}{2}a^2} = \sum_{n=0}^\infty \frac{a^n}{n!} h_n \left( \int_0^T g_t dB_t \right)$$

David Nualart (Kansas University)

Sept. 2016 40/45

イロン 不得 とくほ とくほ とうほ

(ii) On the other hand, using Itô's formula, we obtain

$$M_{T} = 1 + \int_{0}^{T} aM_{s}g_{s}dB_{s}$$
  
= 1 + al\_{1}(g) + a^{2} \int\_{0}^{T} g\_{s} \int\_{0}^{s} M\_{v}g\_{v}dB\_{v}  
= 1 + al\_{1}(g) + a^{2}  $\int_{0}^{T} g_{s} \int_{0}^{s} g_{v}dB_{v} + a^{3} \int_{0}^{T} g_{s} \int_{0}^{s} M_{v}g_{v}dB_{v}$   
=  $\sum_{n=0}^{\infty} \frac{a^{n}}{n!} l_{n}(g^{\otimes n}).$ 

David Nualart (Kansas University)

≤ ≥ Sept. 2016 41/45

▲口 → ▲圖 → ▲ 国 → ▲ 国 →

### Product formula

Let  $f \in L^2_s([0, T]^n)$ , and  $g \in L^2_s([0, T]^m)$ . For any  $r = 0, ..., n \land m$ , we define the *contraction* of *f* and *g* of order *r* to be the element of  $L^2([0, T]^{n+m-2r})$  defined by

$$(f \otimes_r g)(t_1, \ldots, t_{n-r}, s_1, \ldots, s_{m-r}) = \int_{[0,T]^r} f(t_1, \ldots, t_{n-r}, x_1, \ldots, x_r) g(s_1, \ldots, s_{m-r}, x_1, \ldots, x_r) dx_1 \cdots dx_r.$$

イロト 不得 とくほ とくほとう

# Product formula

Let  $f \in L^2_s([0, T]^n)$ , and  $g \in L^2_s([0, T]^m)$ . For any  $r = 0, ..., n \land m$ , we define the *contraction* of *f* and *g* of order *r* to be the element of  $L^2([0, T]^{n+m-2r})$  defined by

$$(f \otimes_r g)(t_1, \ldots, t_{n-r}, s_1, \ldots, s_{m-r}) = \int_{[0,T]^r} f(t_1, \ldots, t_{n-r}, x_1, \ldots, x_r) g(s_1, \ldots, s_{m-r}, x_1, \ldots, x_r) dx_1 \cdots dx_r.$$

- We denote by  $f \otimes_r g$  the symmetrization of  $f \otimes_r g$ .
- Product of two multiple stochastic integrals

$$I_n(f)I_m(g) = \sum_{r=0}^{n \wedge m} r! \binom{n}{r} \binom{m}{r} I_{n+m-2r}(f \otimes_r g).$$

# Wiener Chaos expansion

### Theorem

 $F \in L^2(\Omega)$  can be uniquely expanded into a sum of multiple stochastic integrals :

$$F = E[F] + \sum_{n=1}^{\infty} I_n(f_n).$$

# Wiener Chaos expansion

#### Theorem

 $F \in L^2(\Omega)$  can be uniquely expanded into a sum of multiple stochastic integrals :

$$F=E[F]+\sum_{n=1}^{\infty}I_n(f_n).$$

 For any n ≥ 1 we denote by H<sub>n</sub> the closed subspace of L<sup>2</sup>(Ω) formed by all multiple stochastic integrals of order n. For n = 0, H<sub>0</sub> is the space of constants. Then, we have the orthogonal decomposition

$$L^2(\Omega) = \oplus_{n=0}^{\infty} \mathcal{H}_n.$$

• The theorem follows from the fact that if a random variable  $G \in L^2(\Omega)$  is orthogonal to  $\bigoplus_{n=0}^{\infty} \mathcal{H}_n$ , then it is orthogonal to all random variables of the form  $\left(\int_0^T g_t dW_t\right)^k$ , where  $g \in L^2([0, T])$ ,  $k \ge 0$ . This implies that G is orthogonal to all the exponentials  $\exp\left(\int_0^T g_t dW_t\right)$ , which form a total set in  $L^2(\Omega)$ . So G = 0.

# Integral representation theorem

#### Theorem

Given  $F \in L^2(\Omega, \mathcal{F}_T, P)$  there exists a unique process u in the space  $L^2_T(\mathcal{P})$  such that

$$F=E[F]+\int_0^t u_t dB_t.$$

*Example :*  $F = B_T^3$ . By Itô's formula and integrating by parts

$$B_{T}^{3} = \int_{0}^{T} 3B_{t}^{2} dB_{t} + 3 \int_{0}^{T} B_{t} dt = \int_{0}^{T} 3B_{t}^{2} dB_{t} + 3 \left( TB_{T} - \int_{0}^{T} t dB_{t} \right)$$
  
$$= \int_{0}^{T} 3B_{t}^{2} dB_{t} + 3 \int_{0}^{T} (T - t) dB_{t}$$
  
$$= \int_{0}^{T} 3 \left[ B_{t}^{2} + (T - t) \right] dB_{t}.$$

ヘロト ヘワト ヘビト ヘビト

### Proof :

We know that

$$F=E[F]+\sum_{n=0}^{\infty}I_n(f_n).$$

Then, it suffices to write, for each  $n \ge 1$ ,

$$I_n(f_n) = n! \int_0^T u_n(t) dB_t,$$

where

$$u_n(t) = \int_0^t \int_0^{t_{n-1}} \cdots \int_0^{t_1} f(t, t_1, t_n, \dots, t_{n-1}) dB_{t_1} \cdots dB_{t_{n-1}},$$
  
and take  $u_t = \sum_{n=1}^\infty u_n(t)$ .  $\Box$